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A meta-analysis on verbal working memory in children and adolescents with ADHD

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ABSTRACT

Objective: Previous meta-analyses have demonstrated verbal working memory (WM) dysfunction in children and adolescents with attention deficit/hyperactivity disorder (ADHD); however, the findings are inconsistent. The main objective of this meta-analysis was to investigate the performance of children and adolescents with ADHD in the Digit Span Backwards (DSB) subtest from the Wechsler Intelligence Scale for Children or Wechsler Adult Intelligence Scale. We also sought to provide an updated meta-analysis on WM in children and adolescents with ADHD.

Method: PubMed, PsylNFO, Scopus, and Web of Science were searched to locate studies published between 1990 and 2016 that report DSB scores both of children and adolescents with ADHD and matched controls. Potential moderator variables were also analyzed.

Results: Forty-nine studies comparing children and adolescents with ADHD ($n = 4956$) against healthy controls ($n = 3249$) generated a medium-sized effect (Hedges' g) of 0.56 (95% CI [0.49, 0.64]), indicating poorer verbal WM performance in those with ADHD. A subgroup meta-analysis of studies with participants aged 8–16 years only demonstrated low heterogeneity ($I^2 = 17.06$, $cf.$ 55.50 for the main analysis). Moreover, the meta-regression showed a negative association ($\beta = -.05$, $p = .02$) between DSB performance differences and increasing age, indicating that for every one year increase in age the effect size decreased by .05.

Conclusion: These results, which emanated from the largest meta-analysis concerning verbal WM in ADHD reported to date, reinforce WM as a key domain of cognitive dysfunction in ADHD, and point to age as the main variable influencing DSB performance difficulties.

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Attention-deficit/hyperactivity disorder; digit span backwards; WISC; WAIS; executive function

Introduction

Attention-deficit/hyperactivity disorder (ADHD) is one of the most common neurodevelopmental disorders, with worldwide prevalence estimated at 7.2% in childhood and adolescence (Thomas, Sanders, Doust, Beller, & Glasziou, 2015). The core features of

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ADHD are inattention and/or hyperactivity-impulsivity, along with other symptoms such as academic underachievement, impairment in social functioning and frequent disorganization (American Psychiatric Association, 2013). ADHD symptoms are clustered in three different subtypes of presentation: inattentive (ADHD-I), hyperactive/impulsive (ADHD-HI), and combined (ADHD-C). Inattentive behavior in ADHD includes avoidance or dislike of tasks that require sustained attention (e.g. schoolwork or listening to lectures), whereas hyperactivity/impulsivity encompasses excessive motor movements (e.g. squirming in seat, restlessness) and self-control difficulties (e.g. excessive talking) (American Psychiatric Association, 2013).

Several theoretical models have attempted to shed light on the neurocognitive mechanisms underlying ADHD symptoms. According to Barkley's prominent model, the key dysfunction in ADHD is response inhibition, which enables the ability to suppress a prepotent (dominant) response, to interrupt an ongoing response in favor of a more adaptive response, and to resist internal or external interferences to achieve goal-directed behaviors (Barkley, 1997). Barkley posited that inefficient response inhibition leads to deficiencies in executive functions such as self-regulation (e.g. the ability to control one's emotions) and working memory (WM) (see also Alderson et al., 2017; Sonuga-Barke, 2003, 2005). This chain of disruptions in executive functioning eventually results in the symptoms observed in ADHD (see also Sonuga-Barke, 2003, 2005). For example, inhibitory control difficulties can lead to the child struggling to maintain focus on goal-directed behaviors when distractions from the environment are present, resulting in hyperactivity/impulsivity/inattentiveness symptoms.

In contrast to inhibition models of ADHD, WM models propose that dysfunctional WM, which plays a fundamental role in the development of organized behaviors, is the central characteristic of ADHD (Kofler, Sarver, Harmon, et al., 2018; Rapport, Chung, Shore, & Isaacs, 2001). According to WM models of ADHD, impairments in response inhibition are a byproduct of WM deficits, as the activation of information in WM precedes the decision to suspend or delay a behavioral response (Kofler, Rapport, Bolden, & Altro, 2008). Moreover, it is deficiencies in WM that lead to behaviors characterized as inattentive, impulsive and/or hyperactive, for example in relation to shifts of attentional focus from monotonous or complex tasks to other environmental or internal stimuli (Kofler et al., 2011; Rapport et al., 2009, 2001).

The Rapport et al. (2001) WM model of ADHD was based on assumptions initially formulated by Baddeley and Hitch (1974) regarding how WM works. Baddeley and Hitch (1974) defined WM as a multicomponent system that requires information to be both stored and simultaneously manipulated in the mind to perform complex tasks. According to this approach, WM is composed of a central executive system and two subsidiary short-term systems—the phonological loop and the visuospatial sketchpad. While the former is responsible for storing speech-based information, the latter stores visuospatial information. The central executive, in addition to coordinating its two subsidiary systems, works as a mechanism for controlling attention and monitoring information, and is activated by complex activities that require focused attention (Baddeley, 1992, 1996; Baddeley & Hitch, 1974; Repovš & Baddeley, 2006). More recently, a fourth element was added to this WM model: the episodic buffer, defined as temporary storage responsible for binding information from different sources (such

as phonological loop, visuospatial sketchpad, and long-term memory) and modalities (e.g. visual and verbal) into an organized structure that enables the planning of future actions (Baddeley, 2000; Repovš & Baddeley, 2006).

Setting aside the disagreement regarding whether WM difficulties transpire first, these models of ADHD all postulate that WM difficulties lead to symptoms of ADHD. Moreover, recent research has demonstrated links between WM and abilities fundamental to success in school and in everyday life, such as fluid intelligence (Brydges, Ozolnieks, & Roberts, 2017, in press), reading comprehension (Friedman, Rapport, Raiker, Orban, & Eckrich, 2017), written expression (Eckrich, Rapport, Calub, & Friedman, 2018), mathematical skills (Friedman, Rapport, Orban, Eckrich, & Calub, 2018), and organizational skills (Kofler, Sarver, Harmon, et al., 2018). These findings highlight the importance of having high functioning WM systems.

Evidence suggests that WM deficits may depend on the ADHD subtype. Although not always the case (e.g. Saydam et al., 2015), children with ADHD-I tend to show poorer performance on WM tasks relative to children with ADHD-HI or ADHD-C (Colbert & Bo, 2017; see also Martinussen & Tannock, 2006; Messina, Tiedemann, de Andrade, & Primi, 2006), whereas children with ADHD-C exhibit more difficulties during tasks assessing response inhibition (Desman, Petermann, & Hampel, 2008; Houghton et al., 1999; Lockwood, Marcotte, & Stern, 2001; Romero-Ayuso, Maestú, González-Marqués, Romo-Barrientos, & Andrade, 2006). These results support the model proposed by Diamond (2005) that response inhibition difficulties are the core deficit in ADHD with hyperactivity (ADHD-HI and ADHD-C), whereas WM dysfunction is the core deficit in ADHD-I. In the current quantitative synthesis, we performed a subgroup meta-analysis to determine influences of ADHD subtype on verbal WM function.

Several studies have demonstrated that impairment of verbal WM in ADHD is less evident across empirical findings compared to visuospatial WM (Brocki, Randall, Bohlin, & Kerns, 2008; McInnes, Humphries, Hogg-Johnson, & Tannock, 2003; Rapport et al., 2008; Sowerby, Seal, & Tripp, 2011). For example, Sowerby et al. (2011) found that in children with ADHD, after controlling for comorbid learning and language difficulties, visuospatial WM remained significantly impaired, but impairment of verbal WM remained only marginally significant. However, others have reported that children with ADHD show impairments, albeit not necessarily to the same extent, in both verbal and visuospatial WM after controlling for comorbid learning and language difficulties (e.g. Martinussen & Tannock, 2006; see also Rapport et al., 2008). Moreover, evidence suggests that verbal WM should be impaired. First, poor verbal WM and poor vocabulary skills have been postulated to be intertwined and associated with ADHD symptoms in preschool children, with early verbal WM intervention being posited as an important step in treating ADHD during childhood (Gremillion & Martel, 2012; Gremillion, Smith, & Martel, 2018). Second, in contrast to several empirical findings, previous meta-analyses have identified impairment in verbal WM in children and adolescents with ADHD, but the magnitude of the effect sizes varies according to the number of studies included in each meta-analysis (see Table 1); thus, it was anticipated that a larger meta-analysis may produce a more reliable effect size for verbal WM in ADHD.

A prominent previous meta-analysis investigated several executive functions in ADHD (e.g. planning, inhibition, set-shifting) and estimated an effect size of 0.55

Table 1. Effect sizes from meta-analyses on verbal working memory in attention-deficit/hyperactivity disorder.

Authors (Year)	<i>N</i> (Studies)	<i>N</i> (ADHD)	<i>N</i> (Control)	Effect size	Lower limit	Upper limit
Previous meta-analyses						
Kasper et al. (2012) ^a	34	988	1086	0.69	0.53	0.84
Martinussen et al. (2005) ^{b,c}	13	475	557	0.56	0.29	0.83
Martinussen et al. (2005) ^{b,c}	12	464	538	0.43	0.24	0.62
Willcutt, Doyle, et al. (2005) ^d	11	661	718	0.55	0.44	0.66
Current meta-analysis ^e	49	4956	3249	0.56	0.49	0.64

Note. In this table, only Martinussen et al. (2005) used Cohen's *d*; the other authors chose Hedges' *g* as the effect size. The present meta-analysis estimated Hedges' *g* at 0.56, which is equivalent to Cohens' *d* at 0.56. As the sample size increases, the two effect sizes tend to approach each other in value, thus, avoiding erroneous conclusions when they are compared.

^aKasper et al. (2012) included several verbal WM tasks such as digit span backwards (DSB), number/letter sequencing, 1-back or 2-back, final word recall, and counting span.

^bThe authors reported two results: an outlier was excluded from the second (leaving 12 studies).

^cMartinussen et al. (2005) included several verbal WM tasks such as DSB, number recall of Kaufman Assessment Battery for Children (K-ABC), word order, word span, and digit recall.

^dWillcutt, Doyle, et al. (2005) reported effect sizes based on only two tests (working memory sentence span and DSB).

^ePitzianti et al. (2016) proved to be an outlier and thus was excluded.

ADHD: Attention-Deficit/Hyperactivity Disorder.

regarding verbal WM (Willcutt, Doyle, Nigg, Faraone, & Pennington, 2005). However, the authors did not interpret their results as evidence of verbal WM dysfunction due to small sample sizes and other inconsistencies observed in some of the studies included in the meta-analysis (e.g. sampling procedures and diagnostic criteria). Given that the authors found only a handful of studies that investigated executive functions in ADHD, including verbal WM ($n = 11$), we anticipated that the current meta-analysis with a larger number of studies would help clarify whether verbal WM dysfunction occurs in ADHD. Another important meta-analytical study focused exclusively on WM in children and adolescents with ADHD (Martinussen, Hayden, Hogg-Johnson, & Tannock, 2005), and grouped several WM tasks according to modality (verbal or spatial) and type of processing (storage vs. storage/manipulation). The verbal WM effect size was initially estimated at 0.56, but was adjusted to 0.43 after the removal of an outlier (see Table 1).

Kasper, Alderson, and Hudec (2012) performed a larger meta-analysis on children and adolescents with ADHD, and the effect size was estimated at 0.69 for verbal WM, a larger effect than those estimated by previous quantitative syntheses (see Table 1). This meta-analysis focused on a wide range of potential moderating variables (e.g. age, sex, and number of experimental trials) to explore the effect size variance within and between studies. Interestingly, the authors decided to include only studies with participants between 8 and 16 years of age, a criterion that may have contributed to the lack of influence of age on the effect size variance. In the present study, we included children and adolescents aged 6 to 18 years in the main analysis. In a second step, to enable comparison with the largest meta-analysis to date, the current study first analyzed the age range of 8–16 years as a categorical moderating variable (subgroup meta-analysis), and then conducted a meta-regression with all the ages of ADHD participants to investigate and elucidate whether age influences the effect size variance.

The main goal of this meta-analysis was to investigate verbal WM in children and adolescents with ADHD in comparison with healthy controls, whose scores were

obtained exclusively from the Digit Span Backwards (DSB) subtest of the Wechsler Intelligence Scale for Children (WISC) or the Wechsler Adult Intelligence Scale (WAIS) as these scales are some of the most widely used evaluation tools for neuropsychological assessment in places such as Canada and the United States (Rabin, Barr, & Burton, 2005), Brazil (Ramos & Hamdan, 2016), and Europe (Evers et al., 2012). DSB in particular was chosen because it is one of the most frequently used tests for assessing verbal WM (Conway et al., 2005). In contrast to Digit Span Forwards (DSF), which requires only the temporary storage and maintenance of information in mind, DSB requires storage, maintenance and manipulation of information and thus qualifies as a WM task (reviewed recently in Hurley & Machado, 2018). For instance, empirical findings have indicated that DSB requires executive functioning (Alloway, Gathercole, Kirkwood, & Elliott, 2009; Carlesimo, Fadda, Lorusso, & Caltagirone, 2009; Conklin, Curtis, Katsanis, & Iacono, 2000; Grégoire & Van Der Linden, 1997; Hale, Hoepfner, & Fiorello, 2002; Wells, Kofler, Soto, Schaefer, & Sarver, 2018). Moreover, brain-based evidence shows that DSB enlists brain areas commonly involved in executive functions (e.g. dorsolateral prefrontal cortex and anterior cingulate cortex), which is less evident with DSF (Gerton et al., 2004; Hoshi et al., 2000; Sun et al., 2005; Tian et al., 2014; Yang et al., 2015). For example, activation of dorsal anterior cingulate cortex showed a positive DSB correlation and negative DSF correlation (Yang et al., 2015). Consistent with recommendations (Reynolds, 1997), composite digit span scores were not included in the current meta-analysis because they combine DSF and DSB scores, thus, preventing attribution of performance differences specifically to WM.

The present meta-analytical review adds to the existing literature by providing an updated meta-analysis on verbal WM in children and adolescents with ADHD, presenting 34 studies not previously included in meta-analyses on the topic. The a priori hypothesis was that a meta-analysis carried out using only DSB scores from WISC or WAIS would elucidate whether verbal WM dysfunction is a core neurocognitive feature of ADHD. Finally, the analysis of several potential moderator variables (covariates) was performed to investigate whether they account for some of the effect size variability within and between the studies included in the quantitative synthesis (Holmbeck, 1997).

Method

The systematic review followed the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines (Moher, Liberati, Tetzlaff, Altman, & Group, 2009). The PubMed, PsycINFO, Scopus, and Web of Science databases were searched on 02 January 2017, using the following keywords: (“ADHD” OR “ADD” OR “hyperkinetic syndrome” OR “attention deficit disorder” OR “attention deficit hyperactivity disorder”) AND (“working memory” OR “digit span” OR “digit backward”; see Supporting Information A for more details). Note that the search term “digit span” returned articles that used terms that included these words (e.g. digit span backward and backward digit span). The period of interest covers the years from 1990 to 2016. In addition to the databases, the reference lists of the articles selected during the systematic review were thoroughly examined to find other possible studies of interest. All articles found were transferred to EndNote, after which multiple occurrences of the

same reference were reduced to one. The abstract of each article was then examined to select articles deemed relevant to be included in the meta-analysis.

Included in the search were scientific articles published in English, Spanish, and Portuguese, with the goal of reducing the risk of publication and language biases (Egger, Zellweger-Zähner, et al., 1997). The samples were limited to populations consisting of children (6 years of age and older) and/or adolescents up to and including 18 years of age. Regarding ADHD diagnosis, studies had to report on children and adolescents who had been diagnosed according to the Diagnostic and Statistical Manual of Mental Disorders (DSM-III-R, DSM-IV, DSM-IV-TR, or DSM-V) or the International Statistical Classification of Diseases and Related Health Problems (10th Edition; ICD-10) to be included in the meta-analysis; therefore, an indication of ADHD symptoms without a final diagnosis of ADHD was deemed an insufficient inclusion criterion. Only studies that reported the mean and the *SD* (or *SE* or confidence intervals) of the DSB scores from the WISC or WAIS were included. For studies that reported *SE* or confidence interval, we first converted these statistics into *SD* before conducting the meta-analysis. Following the same approach from previous meta-analyses (Kasper et al., 2012; Martinussen et al., 2005; Willcutt, Doyle, et al., 2005), we included DSB scores from any WISC/WAIS edition (e.g. WISC-III or WISC-R).

Excluded from the meta-analysis were unpublished works, Master's theses, and Doctoral dissertations. Also excluded were studies without a healthy control group, those with pharmacological intervention, and those that included children or adolescents with ADHD reported to exhibit comorbidities such as learning disabilities, schizophrenia, autism, bipolar disorder, or other neurological conditions. However, due to its high incidence in ADHD (Busch et al., 2002), samples with comorbidities of oppositional defiant disorder, conduct disorder, and anxiety disorder were not excluded.

A flow diagram illustrating the systematic review is given in Figure 1, showing that of the 126 articles initially selected, 76 were eliminated for the following reasons: lack of diagnostic criteria necessary for including participants in the ADHD group ($n = 1$); scores from DSB were not reported in the study ($n = 58$); the reported scores from DSB were different from the mean and the *SD* (or *SE* or confidence interval), such as correlation and linear regression ($n = 7$); scores from DSB refer to a previous study conducted by the authors ($n = 1$); studies that reported only the mean from DSB without *SD*, confidence interval, or *SE* ($n = 2$); inclusion of participants younger than 6 or older than 18 years ($n = 6$); lack of information concerning the Wechsler scale used ($n = 1$). Finally, 50 studies were retained and formally included in the meta-analysis; however, the study by Pitzianti et al. (2016) was subsequently eliminated, since it was deemed as an outlier according to Grubbs' test (Grubbs, 1950).

Data extraction

Regarding the data characteristics obtained from the studies, the following variables were compiled: primary author, publication date, title of the article, geographical location, and the version of the Wechsler scale (edition and language). Regarding the sample characteristics, the sample size, age range, IQ, sex, ADHD subtype, diagnostic criteria, and time elapsed since the last use of psychostimulant medication were

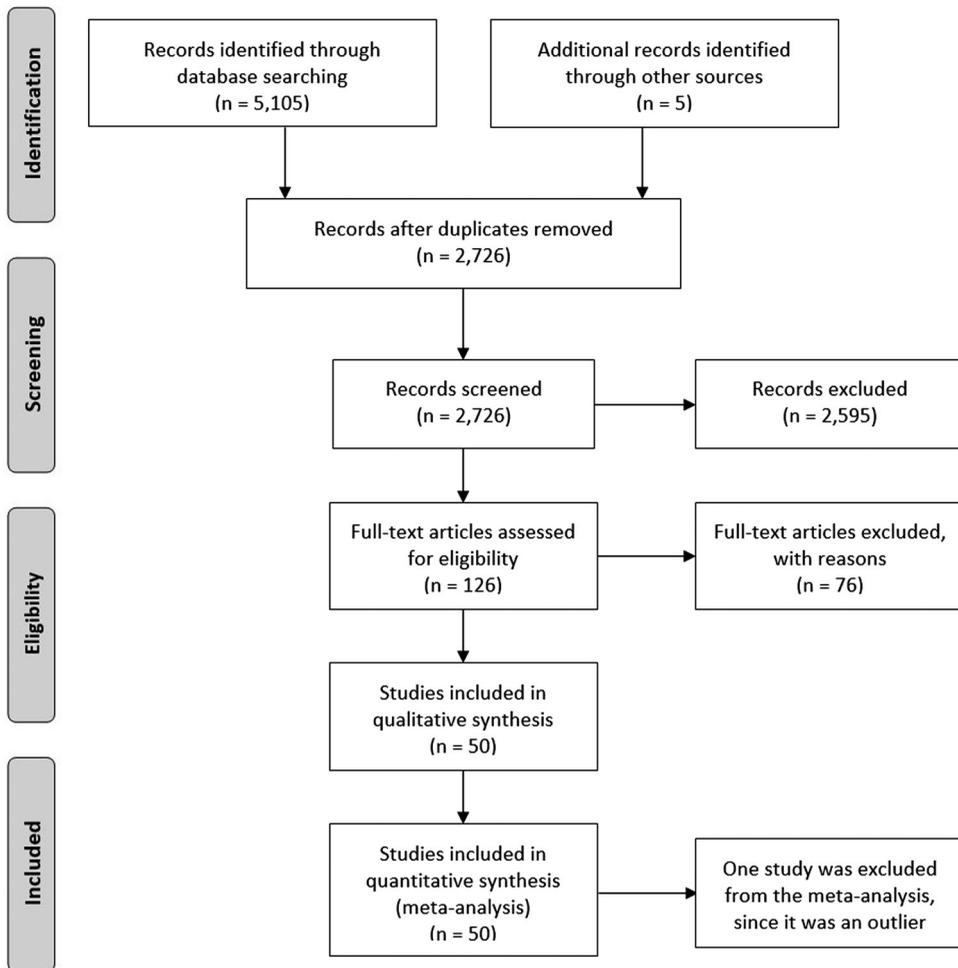


Figure 1. Systematic review flow diagram. Adapted from “Preferred Reporting Items for Systematic Reviews and Meta-Analyses: The PRISMA Statement” by D. Moher, A. Liberati, J. Tetzlaff, D. G. Altman, and The PRISMA Group, 2009. *PLoS Medicine*, 6(7), e1000097, p. 3.

recorded. For studies aimed at genetic investigation, only the scores of the control group and those of the probands were included in the meta-analysis. Table S1 in Supporting Information B shows the moderator variables selected during the systematic review to conduct the subgroup meta-analyses. Considering several studies did not report IQ scores (see Table 2), we did not analyze IQ as a covariate in the meta-regression to avoid confounding findings that may limit generalizability.

Analysis of the data and effect size

A meta-analysis was carried out in R version 3.5.0 (R Core Team, 2018) with RStudio (RStudio Team, 2015), using the meta package (Schwarzer, 2007). The random-effects model is the most appropriate model for the meta-analysis calculation as it is possible that the individual studies measure distinct effects (Engels, Schmid, Terrin, Olkin, & Lau, 2000). The restricted maximum likelihood method was chosen to estimate the

Table 2. Descriptive data and effect sizes from included studies in the meta-analysis.

Studies (Year of publication)	ADHD			Control			Hedges' <i>g</i>	95% CI	
	<i>N</i>	Age	IQ	<i>N</i>	Age	IQ		LL	UL
Aly, Effat, Azb, and Abd Elsamei (2015)	30	10.2 ^a	105.0	30	10.2 ^a	107.9	1.333	0.770	1.896
Banaschewski et al. (2012)	302 ^b	11.3	102.8	259 ^b	12.2	107.6	0.527	0.358	0.696
Casas, Alba, Taverner, Rosello, and Mulas (2006)	33	10.0	106.1	15	10.9	117.8	1.855	1.132	2.578
Casas, Andrés, Castellar, Miranda, and Diago (2011)	21	13.9	108.1	21	13.9	105.8	0.934	0.294	1.575
Chiang and Gau (2008) – ADHD-C	52	8.5	105.0	52	8.7	114.3	0.350	-.038	0.737
Chiang and Gau (2008) – ADHD-I	17	8.5	108.6	52	8.7	114.3	0.346	-.205	0.897
Coutinho, Mattos, and Malloy-Diniz (2009)	186	11.5	96.8	80	12.4	98.5	0.417	0.153	0.682
Fan, Gau, and Chou (2014)	25	10.9	107.2	23	11.2	109.1	0.601	.021	1.181
Gau and Chiang (2013)	389 ^c	12.0 ^a	–	317 ^c	12.0 ^a	–	0.556	0.405	0.707
Gau and Shang (2010)	279	12.5	103.0	173	12.6	110.7	0.478	0.286	0.671
Gau, Chiu, Shang, Cheng, and Soong (2009)	53	12.7	107.3	53	12.7	108.6	0.245	-.137	0.627
Gremillion and Martel (2012)	201 ^b	9.7	100.8	154 ^b	9.8	104.7	0.395	0.183	0.607
Healey and Rucklidge (2006)	29	11.4	106.0	30	11.1	117.1	0.556	.035	1.077
Karalunas and Huang-Pollock (2013)	91	10.2	104.4	62	10.6	106.9	0.299	-.026	0.623
Kaufmann and Nuerk (2006)	16	10.2	–	16	10.4	–	-.158	-.852	0.536
Kim et al. (2014)	17	10.1	–	17	10.2	–	0.769	.069	1.469
Lee et al. (2008) – ADHD-C	200	8.6	102.3	47	10.2	119.0	0.841	0.515	1.168
Lee et al. (2008) – ADHD-HI	60	8.3	105.9	47	10.2	119.0	0.809	0.412	1.207
Lee et al. (2008) – ADHD-I	113	9.1	102.5	47	10.2	119.0	0.907	0.553	1.262
Lee et al. (2008) – HKD	46	8.0	103.9	47	10.2	119.0	1.073	0.637	1.509
Manassis, Tannock, Young, and John (2007) – ADHD	21	9.6	103.4	35	9.7	110.4	0.753	0.193	1.313
Manassis, Tannock, Young, and John (2007) – ADHD + ANX	35	9.8	103.5	35	9.7	110.4	0.457	-.018	0.932
Martinussen and Tannock (2006)	60 ^d	8.7	101.7	25 ^d	9.1	115.6	0.587	0.111	1.062
McAuley et al. (2014)	64	13.7	102.8	28	13.6	117.5	-.122	-.566	0.323
McInerney and Kerns (2003)	30	10.1	–	30	10.1	–	0.786	0.259	1.312
Mello et al. (2013)	12	10.7	–	11	9.1	–	1.132	0.238	2.026
Nguyen et al. (2014)	77 ^b	11.5	91.6	141 ^b	12.4	104.3	0.387	0.107	0.667
Nyman et al. (2010)	30	8.7	94.7	30	8.6	108.3	0.731	0.207	1.896
O'Brien, Dowell, Mostofsky, Denckla, and Mahone (2010) – ADHD Boys	30	10.5	107.6	48	10.1	115.3	0.290	-.169	0.927
O'Brien, Dowell, Mostofsky, Denckla, and Mahone (2010) – ADHD Girls	26	9.8	109.0	42	10.4	113.0	0.433	-.062	0.595
Øie, Sundet, and Rund (1999)	20	14.1	98.6	30	15.7	110.5	0.475	-.099	1.575
Pasini, Paloscia, Alessandrelli, Porfirio, and Curatolo (2007) – ADHD-C	25	10.1	–	44	10.6	100.0	0.711	0.205	0.737
Pasini, Paloscia, Alessandrelli, Porfirio, and Curatolo (2007) – ADHD-I	25	11.4	–	44	10.6	100.0	0.682	0.177	0.897
Passolunghi, Marzocchi, and Fiorillo (2005)	10	9.8	105.1	10	9.9	107.5	1.019	.074	0.682
Pitzianti et al. (2016)	13	10.2	97.5	13	11.9	106.3	3.059	1.873	4.244
Qian, Shuai, Cao, Chan, and Wang (2010) – ADHD	89	9.1	108.2	116	9.2	112.8	0.185	-.092	0.462
Qian, Shuai, Cao, Chan, and Wang (2010) – ADHD + ODD	53	9.3	107.0	116	9.2	112.8	0.249	-.077	0.575
Qian, Shuai, Chan, Qian, and Wang (2013) – 7-to 8-year old	153	7.8	104.6	55	8.1	111.9	0.276	-.033	0.585
Qian, Shuai, Chan, Qian, and Wang (2013) – 9-to 10-year old	182	9.3	105.9	100	9.5	116.3	0.293	.048	0.539
Qian, Shuai, Chan, Qian, and Wang (2013) – 11- to 12-year old	76	11.4	105.9	62	11.3	115.4	0.481	0.141	0.822
Qian, Shuai, Chan, Qian, and Wang (2013) – 13- to 15-year old	104	13.7	106.4	32	13.5	117.6	0.199	-.198	0.596
Rodríguez Pérez et al. (2009)	40	11.4	100.4	124	10.9	109.6	0.873	0.504	1.242
Rosenthal et al. (2006) – ADHD-C	28	11.4	102.0	27	11.5	109.5	0.693	0.147	1.238
Rosenthal et al. (2006) – ADHD-I	12	12.4	95.8	27	11.5	109.5	-.301	-.984	0.383
Rucklidge and Tannock (2002)	35	15.2	102.2	37	14.9	111.0	0.253	-.211	0.718
Saydam et al. – ADHD-C	37	8.9	103.6	36	9.3	108.3	0.516	.049	0.982

(continued)

Table 2. Continued.

Studies (Year of publication)	ADHD			Control			Hedges' <i>g</i>	95% CI	
	<i>N</i>	Age	IQ	<i>N</i>	Age	IQ		LL	UL
Saydam et al. – ADHD-C + ODD/CD	37	9.1	102.3	36	9.3	108.3	0.596	0.126	1.065
Saydam et al. – ADHD-I	37	9.9	103.6	36	9.3	108.3	-.158	-.618	0.301
Schachar et al. (2007) – ADHD-C	353	8.8	100.6	67	9.8	118.4	0.826	0.559	1.094
Schachar et al. (2007) – ADHD-HI	142	8.6	104.8	67	9.8	118.4	0.773	0.473	1.073
Schachar et al. (2007) – ADHD-I	237	9.3	101.7	67	9.8	118.4	0.821	0.542	1.100
Schachar et al. (2007) – HKD	72	8.2	101.4	67	9.8	118.4	1.018	0.664	1.372
Siklos and Kerns (2004)	19	10.1	71.3	19	10.0	71.4	0.814	0.149	1.479
Simone, Bédard, Marks, and Halperin (2016)	63	8.6	103.6	51	8.5	110.4	0.573	0.196	0.950
Skowronek, Leichtman, and Pillemer (2008)	12	12.2	–	17	11.5	–	0.949	0.165	1.733
Soraa, Iraola, Balluerka, and Soraa (2009)	12	8.5	–	13	9.0	–	1.055	0.208	1.902
Sowerby et al. (2011)	40	8.4	94.7	40	8.4	101.4	0.371	-.071	0.814
Spronk et al. (2013)	15	14.8	99.1	19	14.8	96.5	.000	-.677	0.677
Tiffin-Richards, Hasselhorn, Woerner, Rothenberger, and Banaschewski (2008)	20	11.6	97.4	19	11.7	107.9	0.852	0.193	1.511
Tillman, Bohlin, Sørensen, and Lundervold (2009)	45	10.3	–	120	9.8	–	0.505	0.158	0.852
Toplak and Tannock (2005)	46	15.6	104.1	44	15.3	110.7	0.861	0.428	1.294
Toplak, Rucklidge, Hetherington, John, and Tannock (2003)	35	15.2	102.2	39	15.0	110.1	0.266	-.193	0.724
Udal et al. (2013)	26	12.6	97.8	69	12.3	99.9	0.723	0.260	1.186
Willcutt, Pennington, Olson, Chhabildas, and Hulslander (2005)	113	11.2	104.3	151	11.5	113.6	0.590	0.341	0.839
Wu, Anderson, and Castiello (2002)	58	10.5	103.3	29	10.6	113.7	0.426	-.024	0.877
B. R. Yang et al. (2007)	40	8.5	99.3	40	8.6	108.3	0.723	0.270	1.176
B. R. Yang et al. (2011)	100	8.4	98.5	100	8.5	106.8	1.044	0.749	1.340
Full effect size							0.573	0.495	0.650
Full effect size without Pitzianti et al. (2016)							0.563	0.488	0.638

Note. Dashes indicate information not provided. CI: Confidence Interval. LL: Lower Limit. UL: Upper Limit. ADHD: Attention-Deficit/Hyperactivity Disorder (C: combined subtype; I: inattentive subtype; HI: hyperactivity-impulsivity subtype). HKD: Hyperkinetic Syndrome. ODD: Oppositional-Defiant Disorder. CD: Conduct Disorder. ANX: Anxiety Disorder.

^aMean age reported by the authors refers to the full sample (ADHD and controls).

^bThe number of children and adolescents whose DSB scores were reported in the study. However, the mean and *SD* of age and IQ refer to the full sample in each group (clinical and control).

^cDescriptive statistics reported only for the full sample and not separated according to group (clinical and control).

^dDSB data were not available for some ADHD and control participants; the full sample was composed of 62 participants with ADHD and 34 control participants.

between-study variance (Patterson & Thompson, 1971). Hedges' *g*, also known as the corrected effect size, was used. It is an unbiased estimate of the effect size of the population (Hedges, 1981) and is appropriate for situations involving studies with small sample sizes ($n < 20$) (Hedges & Olkin, 1985; Lakens, 2013). The interpretation of Hedges' *g* is based on benchmarks suggested by Cohen (1988): 0.2 = a small effect; 0.5 = a medium effect; 0.8 = a large effect. The common language effect size (McGraw & Wong, 1992), for which interpretation is based on percentage (0% to 100%), was also used. The calculation of the effect sizes was based on the mean and the *SD* of DSB scores from the samples with ADHD and the control group.

As a mere matter of convention, this meta-analysis uses positive *g* values to indicate effects unfavorable to the group with ADHD. Possible outliers, defined as "data points that lie at a distance from other data points because they are the result of inaccuracies" (Aguinis, Gottfredson, & Joo, 2013, p. 281), were identified using Grubbs' test (Grubbs, 1950). Raw scores and scaled or age-adjusted scores were included in the meta-analysis, which is deemed acceptable since in all cases ADHD scores were

compared against age-matched controls; this rationale is also in accordance with previous meta-analyses (Kasper et al., 2012; Martinussen et al., 2005; Willcutt, Doyle, et al., 2005).

Heterogeneity and publication bias

Heterogeneity occurs when a difference found in the results from the individual studies either is not explained by sample error or exceeds the sample error (Cumming, 2012) and may not be satisfactorily attributed to chance (Engels et al., 2000). In addition to the Q statistic, which measures the effect size variance regarding the degrees of freedom, the interpretation of heterogeneity in this meta-analysis was based on the I^2 statistic (Higgins & Thompson, 2002), which measures the percentage of the effect size variability not caused by sample error and generates an amplitude that varies between 0% and 100% (Higgins & Thompson, 2002; Huedo-Medina, Sánchez-Meca, Marín-Martínez, & Botella, 2006). To interpret I^2 results, the benchmarks proposed by Higgins, Thompson, Deeks, and Altman (2003)—low (25%), moderate (50%), and high (75%)—were adopted in the present meta-analysis.

Publication bias occurs when published studies are not representative of the entire body of research that has been conducted in a specific area of interest, which can result in incorrect conclusions and harmful treatments (Rothstein, Sutton, & Borenstein, 2005). The preference for studies with statistically significant results (the file drawer problem) remains the main cause of publication bias (Button, Bal, Clark, & Shipley, 2016). In this meta-analysis, publication bias was studied principally using funnel plots. When no bias exists, the funnel plot assumes the shape of an inverted funnel, and the plotted data exhibit a symmetrical distribution on both sides (Egger, Smith, Schneider, & Minder, 1997; Sterne & Egger, 2001). In consideration of the fact that a bias analysis-based exclusively on funnel plots is primarily visual (Egger, Smith, et al., 1997), precarious (Ioannidis, 2008) and subjective (Duval, 2005; Sterne, Becker, & Egger, 2005), its use was complemented with Duval and Tweedie's trim-and-fill method (Duval & Tweedie, 2000). This method assumes that a meta-analysis has been affected by publication bias when the funnel plot is asymmetrical. Thus, besides the included samples (n_0), the trim-and-fill method estimates the number of relevant samples (k_0) not included in the meta-analysis due to publication bias. Since k_0 and its respective effect sizes are unknown, the trim-and-fill method estimates these values and plots the absent samples on the funnel plot, thus, making it symmetrical and recalculating the "real" center of the funnel (Duval, 2005; Duval & Tweedie, 2000).

Results

The descriptive data from the samples and estimated effect sizes (Hedges' g) are shown in Table 2. The full effect size was initially moderate (0.57, 95% CI [0.50, 0.65]). However, the study by Pitzianti et al. (2016) with an effect size of 3.06 (95% CI [1.87, 4.24]) was subsequently removed from the meta-analysis since it proved to be an outlier according to the Grubbs test ($G = 5.13$, $p < .05$). The exclusion of this study lowered the heterogeneity percentage (I^2) from 59.58 ($Q = 163.29$, $df = 66$) to 55.50

($Q = 146.07$, $df = 65$). Therefore, the final estimated g value of 0.56 (95% CI [0.49, 0.64]) was adopted for this meta-analysis. After excluding the study by Pitzianti et al. (2016), this meta-analysis included a sample of 4956 participants with ADHD and a sample of 3249 healthy controls. A total of 49 studies were included, and DSB scores were obtained from 66 ADHD samples and 53 control samples.

Four studies were published in Spanish, and the remaining studies in English. All the studies included in the meta-analysis were carried out using the WISC and/or WAIS versions developed for use with the population of the country in which it was conducted. Only one study found no difference between the ADHD group and the control group, that is, $g = 0$ (Spronk, Vogel, & Jonkman, 2013), and four studies reported better performance in the ADHD group (Kaufmann & Nuerk, 2006; McAuley, Crosbie, Charach, & Schachar, 2014; Rosenthal, Riccio, Gsanger, & Jarratt, 2006; Saydam, Ayvasik, & Alyanak, 2015). A Hedges' g value estimated at 0.56 indicates that the average DSB performance of the children and adolescents with ADHD was found to be slightly more than one half of a SD below the average of the healthy controls. According to Dunlap (1999), an effect size of 0.56 corresponds to 0.65 in the common language effect size (the statistic that allows one to determine the real effect size in terms of probability). This means that a child chosen randomly from the control group has a 65% probability of achieving a higher score than a child with ADHD.

The moderate degree of heterogeneity ($I^2 = 55.50$; $Q = 146.07$, $df = 65$) suggests that 55% of the variability between the effect sizes is a consequence of actual heterogeneity and is not explained by sample error (Huedo-Medina et al., 2006). Table 3 summarizes the results from the subgroup meta-analyses, and Table S1 in Supporting Information B shows the details about the moderator variables included in the meta-analysis. Age seemed to be the main explanation for the moderate level of heterogeneity, since the subgroup meta-analysis of 29 studies that included only children and adolescents aged 8 to 16 years estimated a medium-sized effect at 0.45 (95% CI [0.37, 0.53]), with a low level of heterogeneity ($I^2 = 17.06$). To address the question of whether the effect size varied by age, we performed a meta-regression with the average age of the ADHD group as a covariate. Findings suggest an association between age and verbal WM performance differences. Age accounted for 12.76% (R^2) of the variance in true effects, with a medium level of residual heterogeneity ($I^2 = 54.49$; $Q = 136.21$, $df = 64$), which means that approximately 55% of the unaccounted variability in the meta-regression may be explained by other covariates. The coefficient for age ($\beta = -.05$, $p = .02$, 95% CI [-.08, -.01]) suggested that for every increase of one year of age the effect size decreased by .05. Figure 2 shows the dispersion of the studies with ADHD age as a covariate. The size of the symbols in Figure 2 reflects the weight given to each study included in the meta-analysis; larger symbols represent studies with a more precise effect (greater statistical power that is reflected in smaller confidence intervals).

Regarding other moderators (see Table 3), it was initially thought that ADHD subtypes could explain the variability of the data; however, these subgroup meta-analyses showed that only ADHD hyperactive/impulsive (ADHD-HI, $n = 2$) and hyperkinetic syndrome (HKD, $n = 2$) subtypes were homogeneous, with $I^2 = .00$. The edition of the Wechsler scale used also provided minimal insight regarding the variability of the

Table 3. Subgroup meta-analyses.

Moderators	Subgroups	<i>N</i>	Effect size	Lower limit	Upper limit	<i>I</i> ²	<i>Q</i>
Age range	8- to 16-year old	29	0.450	0.370	0.531	17.06	33.76
	Other age ranges	37	0.636	0.526	0.745	65.21	103.49
Sex	Boys	8	0.692	0.278	1.106	68.93	22.53
	Girls	1	0.433	-.062	0.927	–	–
	Mixed	51	0.567	0.481	0.653	57.60	117.92
	Not Informed	6	0.501	0.407	0.596	.00	3.84
ADHD subtype	Mixed	32	0.438	0.368	0.508	32.75	46.10
	ADHD-C	15	0.652	0.520	0.785	45.62	25.74
	Not Informed	8	0.742	0.516	0.967	43.44	12.38
	ADHD-I	7	0.493	0.113	0.874	24.70	75.71
	ADHD-HI	2	0.786	0.547	1.026	.00	.02
	HKD	2	1.040	0.765	1.315	.00	.04
Wechsler edition	WISC-III	36	0.594	0.494	0.694	52.38	73.49
	WISC-R	12	0.691	0.457	0.926	58.76	26.67
	WISC-IV	8	0.405	0.283	0.527	.00	2.72
	C-WISC	8	0.425	0.210	0.641	72.89	25.82
	WISC-III/WAIS-III	2	0.797	0.481	1.113	.00	0.18
Diagnostic criteria	DSM-IV	50	0.535	0.457	0.612	52.57	103.31
	DSM-IV-TR	11	0.624	0.330	0.919	65.07	28.63
	ICD-10	2	1.040	0.765	1.315	.00	.04
	DSM-III-R	1	0.475	-.099	1.049	–	–
	DSM-V	1	0.594	0.217	0.972	–	–
Medication status	24 h	19	0.461	0.371	0.550	18.30	22.03
	48 h	11	0.807	0.684	0.930	23.57	13.08
	Not informed	10	0.707	0.373	1.042	74.29	35.01
	Nonmedicated	8	0.350	0.206	0.494	32.66	10.40
	On the Day	6	0.378	0.128	0.627	41.67	8.57
	Naïve	5	0.718	0.444	0.992	58.54	9.65
	7 half-lives	2	0.581	0.219	0.943	.00	0.63
Sample size	more than 30	41	0.554	0.469	0.639	60.40	101.00
	less than 30	25	0.592	0.431	0.754	46.55	44.90
Wechsler language	Canadian English	17	0.709	0.569	0.850	49.43	31.64
	American English	10	0.438	0.329	0.547	13.86	10.45
	Chinese	10	0.413	0.232	0.593	65.33	25.96
	Chinese (Taiwan)	4	0.505	0.394	0.617	.00	2.40
	Spanish (Spain)	4	1.126	0.696	1.555	48.04	5.77
	Italian	3	0.737	0.403	1.071	.00	0.40
	Turkish	3	0.316	-.155	0.786	67.53	6.16
	Norwegian	3	0.563	0.313	0.813	.00	0.66
	Brazilian Portuguese	2	0.642	-.009	1.292	55.76	2.26
	New Zealander English	2	0.449	0.112	0.786	.00	0.28
	German	2	0.353	-.636	1.343	76.62	4.28
	Dutch	1	.000	-.677	0.677	–	–
	Finnish	1	0.731	0.207	1.255	–	–
	Korean	1	0.769	.069	1.469	–	–

Note. ADHD: Attention-Deficit/Hyperactivity Disorder (C: combined subtype; I: inattentive subtype; HI: hyperactive-impulsive subtype). WISC: Wechsler Intelligence Scale for Children (III: third edition; R: revised; IV: fourth edition). C-WISC: Chinese Revised Wechsler Intelligence Scale for Children. WAIS: Wechsler Adult Intelligence Scale (III: third edition). DSM: Diagnostic and Statistical Manual of Mental Disorders (III-R: third edition, revised; IV: fourth edition; IV-TR: fourth edition, text revised; V: fifth edition). ICD-10: International Statistical Classification of Diseases and Related Health Problems – 10th Edition.

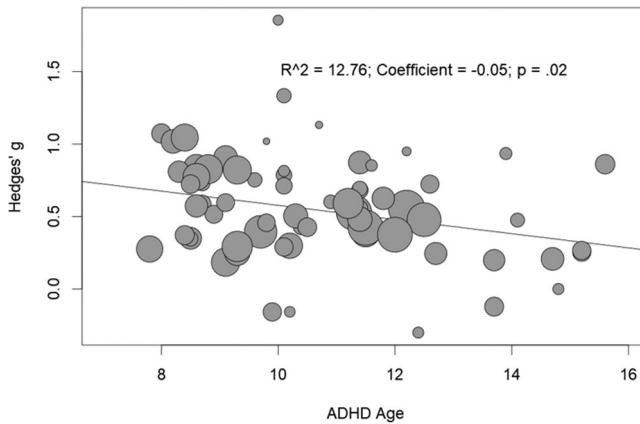


Figure 2. Meta-regression of 66 Attention-Deficit/Hyperactivity Disorder (ADHD) samples showing that age influences the effect size.

data. Only eight studies that applied the WISC-IV exhibited homogeneity ($I^2 = .00$); the remaining editions of the scale all exhibited a moderate to high degree of heterogeneity. To investigate the possible influence of the language of the Wechsler scale on effect size variability, we also performed a subgroup meta-analysis with the language version as a covariate. Interestingly, studies with the American English version ($n = 10$) estimated more homogeneous results ($I^2 = 13.86$) than studies with the Canadian English version ($n = 17$; $I^2 = 49.43$). Ten ADHD samples from China provided medium heterogeneity ($I^2 = 65.33$), but the samples from Taiwan ($n = 4$) were homogeneous ($I^2 = .00$). Homogeneous results ($I^2 = .00$) were also observed for the Italian ($n = 3$), Norwegian ($n = 3$), and New Zealander English ($n = 2$) samples. Note, however, that for most of the language versions, low statistical power due to small sample sizes (two to four studies) limits interpretation of the results.

Regarding the time elapsed since the most recent psychostimulant medication, there was some indication that withdrawing medication decreased heterogeneity. Namely, both subgroups who had their medication withdrawn for at least 1 day exhibited a low percentage of heterogeneity ($I^2 < 24$). In contrast, the nonmedicated subgroup ($I^2 = 32.66$) and the subgroup in which medication was suspended on the day of the test ($I^2 = 41.67$) exhibited moderate degrees of heterogeneity. The hypothesis that the studies including less than 30 participants (in one or both of the groups) are more heterogeneous ($I^2 = 46.55$) than studies containing more than 30 participants ($I^2 = 60.40$) was discarded since the calculated heterogeneity percentages were not in the predicted direction (Higgins et al., 2003). The diagnostic criteria, medication status, and sex of participants also failed to provide plausible explanations.

The analysis of the funnel plot shown in Figure 3, in which the x -axis gives the g values and the y -axis shows the respective SE , indicates a possible publication bias in the main meta-analysis, indicated by the asymmetrical form, which is more heavily weighted toward the left (without the white symbols). If there was homogeneity, at least 95% of the samples would fall within the funnel (Sterne & Egger, 2001). Duval and Tweedie's trim-and-fill method indicates that at least nine statistically nonsignificant samples are necessary to correct the publication bias (represented in Figure 3 by

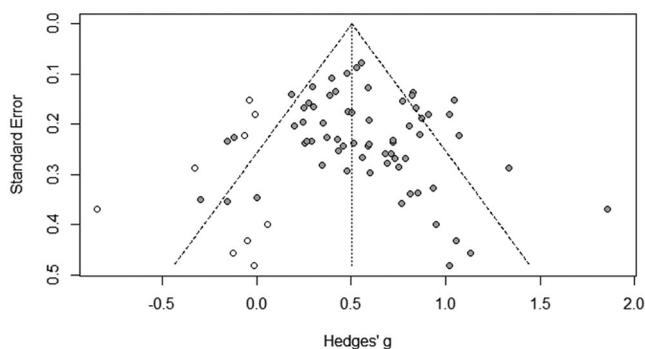


Figure 3. Funnel plot for the analysis of publication bias. Grey symbols indicate the samples ($n = 66$) actually included in the meta-analysis, and white symbols indicate the number of samples with nonsignificant results ($n = 9$) necessary to reduce the publication bias, thus making the funnel symmetrical.

white symbols). After applying Duval and Tweedie's trim-and-fill method, the g value changed to 0.50 (95% CI [0.42, 0.58]). However, it must be noted that the confidence interval greatly overlapped with the full effect size of 0.56 (95% CI [0.49, 0.64]), which means that the bias observed in this meta-analysis is small, and both of the effect sizes fall in a medium-sized range (Cohen, 1988).

Discussion

As far as we know, this is the largest meta-analysis concerning verbal WM in ADHD conducted to date. This meta-analysis makes an important contribution to the existing literature by indicating that children and adolescents with ADHD exhibit poorer DSB performance, with a medium-sized effect of 0.56 (95% CI [0.49, 0.64]). Table 1 shows that this result coincides with those found by Willcutt, Doyle, et al. (2005) and by Martinussen et al. (2005), regarding verbal WM in children and adolescents with ADHD. In contrast, the meta-analysis by Kasper et al. (2012) estimated a slightly larger effect of 0.69. The main explanation for this discrepancy likely relates to the rationale underlying the selection of tasks to estimate a verbal WM index.

Both Willcutt, Doyle, et al. (2005) and Martinussen et al. (2005) selected verbal WM tasks based on the assumption that WM requires storage, maintenance and manipulation of information in mind—a rationale that was also applied in the current meta-analysis. Conversely, Kasper et al. (2012) followed a different approach, in which verbal WM tasks were classified as requiring low (e.g. DSF or number recall) or high (e.g. DSB or 2-back) central executive demand. However, the inclusion of DSF to estimate a single verbal WM index may be a confounding variable since DSF and DSB do not tap the same cognitive processes (Gignac, Kovacs, & Reynolds, 2018; Reynolds, 1997). For instance, DSF does not require mental manipulation of information and thus qualifies as a measure of short-term memory not WM (Diamond, 2013). It seems that the selection of verbal WM tasks may account for the discrepancy in effect sizes between the present meta-analysis and the meta-analysis by Kasper et al. (2012).

A second main finding from the current study relates to age as the strongest covariate that may explain the moderate level of heterogeneity. The meta-analysis of subgroups with children and adolescents aged 8 to 16 years decreased the level of heterogeneity (I^2) from 55.50 (moderate) to 17.06 (low), and the meta-regression suggested that age accounts for 12.76% (R^2) of the variance of the effect size. These findings suggest that poorer verbal WM performance in ADHD is age-related, with younger children with ADHD presenting with greater difficulties, as has been indicated in previous studies (Siegel & Ryan, 1989; Sowerby et al., 2011). This effect of age may be accounted for by two main reasons. First, whereas the ability to hold and manipulate information in mind (as required by the DSB task) normally develops and improves progressively during childhood and adolescence (Cowan, Scott Saults, & Elliott, 2002; Crone, Wendelken, Donohue, van Leijenhorst, & Bunge, 2006; Diamond, 2013; Luciana, Conklin, Hooper, & Yarger, 2005), younger children with ADHD have shown particular problems, which suggests that verbal WM in ADHD develops at a slower pace compared to typical controls (Sowerby et al., 2011). Second, inhibitory control, which enables an individual to maintain focus on a specific task despite internal or external interferences and contributes to the efficacy of verbal WM, has been postulated as a core deficit in ADHD (Barkley, 1997) that may be age sensitive (Macdonald, Beauchamp, Crigan, & Anderson, 2014). In relation to this, in addition to the normal developmental struggle with inhibitory control during childhood, younger children with ADHD also face poorer inhibitory control related to their ADHD (Oosterlaan, Logan, & Sergeant, 1998; Pievsky & McGrath, 2018), thus, compounding difficulties at younger ages.

A third important finding from this quantitative synthesis was the low degree of heterogeneity ($I^2 = 13.86$) for the American English samples in the subgroup meta-analysis of Wechsler scale language version (see Table 3), which indicates that language version may moderate the effect size variance. The idea that language version may be a moderating factor fits with previous research that found differences in digit span performance across four different language versions of WAIS-III (Mexico, Spain, Puerto Rico, and the United States), such that Mexican samples presented with lower digit span scores than Spanish or American normative samples (Funes, Rodriguez, & Lopez, 2016). However, because seven out of the ten American English samples analyzed in the current study were aged 8–16 years, it may be that age played a more prominent role in the low level of heterogeneity than the language version of the scale. The Taiwanese, Italian, New Zealander English, and Norwegian samples showed homogeneous results ($I^2 = .00$), whereas moderate to high heterogeneity (I^2 between 48.04 and 76.62) was estimated for studies from Spain, Turkey, Brazil, and Germany; however, these findings should be interpreted with caution since the analyses included only two to four studies. Jackson and Turner (2017) recently noted that at least five studies are necessary to provide a satisfactory minimum level of statistical power in random-effects models. Thus, while these results provide some indication that language version may moderate the effect size variance; future research will be needed to confirm the results.

The remaining selected moderating variables largely failed to provide evidence explaining the moderate degree of variability in the main analysis ($I^2 = 55.50$). For

example, the subgroup meta-analysis of Wechsler scale edition for the most part also presented moderate degrees of heterogeneity (ranging from 52% to 73%), except for the eight studies with WISC-IV, which provided homogeneous results ($I^2 = .00$). Since previous research indicates that the subtype ADHD-I exhibits a closer relationship with WM functioning compared to the subtypes ADHD-C and ADHD-HI (e.g. Colbert & Bo, 2017), we predicted that the ADHD subtypes might explain some of the effect size variance, but this was not verified by the subgroup meta-analysis. Only seven samples with ADHD were diagnosed exclusively with ADHD-I, compared to 15 samples with ADHD-C and 32 samples with a mixture of two or three subtypes. It may be that the inclusion of a greater number of studies with participants diagnosed specifically with ADHD-I could have explained part of the unexplained variance.

From a clinical standpoint, the results reinforce verbal WM as a key cognitive function affected by ADHD, in accordance with our primary hypothesis. Verbal WM is a pivotal cognitive function in the classroom, given that students rely on it to accomplish daily school activities (e.g. Alloway & Alloway, 2010; Orban, Rapport, Friedman, Eckrich, & Kofler, 2018). Furthermore, children with poorer verbal WM may present with problems such as academic underachievement, high levels of distractibility, poor creativity in complex tasks, difficulties in following instructions, and failures to complete assignments (Alloway et al., 2009; Gathercole et al., 2008). Thus, the assessment of verbal WM may help towards preventing educational underachievement in children with ADHD. In relation to this, although, most past efforts to train WM in children with ADHD have had limited success, a recent approach has demonstrated promise (Kofler, Sarver, Austin, et al., 2018).

An important limitation of this meta-analysis is the reliance on only DSB to provide a verbal WM index in ADHD, which prevents generalizability to other verbal WM tasks such as *n*-back and reading/listening span. Other tasks might provide different effect sizes. Inclusion of a variety of WM tasks, considered separately via subgroup meta-analyses, would enable an investigation of the influence of task complexity on verbal WM performance (e.g. 2-back vs. 3-back). However, care would need to be taken to ensure the tasks are matched regarding aspects other than WM, to avoid misinterpretation. A second limitation of the present meta-analytical study is that the funnel plot analysis and Duval and Tweedie's trim-and-fill method revealed a publication bias in the main meta-analysis. Although there is no consensus among researchers regarding the inclusion of unpublished studies in a meta-analysis (Hopewell, Clarke, & Mallett, 2005), it is probable that the inclusion of unpublished studies with statistically non-significant results would correct the asymmetry of the funnel plot (Egger, Smith, et al., 1997). Conversely, an important characteristic of the current meta-analysis that tends to reduce publication bias is the inclusion of studies published in Spanish, a strength not commonly found in meta-analytic studies. Sterne, Egger, and Smith (2001) pointed out that studies with statistically significant results have more probability of being published in English than those with nonsignificant outcomes; as a consequence, meta-analyses including only studies published in English may suffer from publication and language biases.

In conclusion, this was the first meta-analysis focused exclusively on DSB for estimating effect sizes for verbal WM differences in children and adolescents with

ADHD. The results clearly demonstrate that children and adolescents with ADHD have significantly poorer verbal WM compared to typically developing controls. Additionally, our quantitative synthesis provided an update on previous meta-analyses concerning verbal WM in ADHD by including 34 studies not included in past meta-analyses on the topic (Kasper et al., 2012; Martinussen et al., 2005; Willcutt, Doyle, et al., 2005). The meta-regression reported here suggests an association between age and DSB performance differences, such that verbal WM performance difficulties appear to lessen with increasing age. Future meta-analyses should include scores from DSF and DSB to explore potential performance differences in ADHD between verbal short-term memory (DSF) and verbal WM (DSB). In addition, future meta-analyses should determine the true influence of Wechsler scale language version on verbal WM performance.

Disclosure statement

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